

1. ZERO INFLATED POWER SERIES DISTRIBUTION

A1.1. Let X_1, X_2, \dots, X_m be independently and identically distributed (iid)

random variables (r. v.) with ZIPSD then,

$$\text{i. } E(m_0) = m \left(1 - \pi + \pi \frac{a_0}{f(\theta)} \right).$$

$$\text{ii } E \left(\sum_{i=1}^{\infty} m_i \right) = m \pi \left(1 - \frac{a_0}{f(\theta)} \right).$$

$$\text{iii } E \left(\sum_{i=1}^{\infty} i m_i \right) = \frac{m \pi \theta f'(\theta)}{f(\theta)}.$$

Proof.-

i. Define ,

$$Y_i = \begin{cases} 1, & \text{if } X_i = 0 \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} E(m_0) &= E \left(\sum_{i=1}^m Y_i \right), \\ &= \sum_{i=1}^m E(Y_i) = \sum_{i=1}^m 1 * P(X_i = 0) , \\ &= \sum_{i=1}^m \left(1 - \pi + \pi \frac{a_0}{f(\theta)} \right) = m \left(1 - \pi + \pi \frac{a_0}{f(\theta)} \right), \end{aligned}$$

ii. To find $E\left(\sum_{i=1}^{\infty} m_i\right)$.

$$\text{Define, } Z_{ij} = \begin{cases} 1, & \text{if } X_j = i \\ 0, & \text{otherwise} \end{cases}$$

$$\sum_{j=1}^{\infty} Z_{ij} = m_i = \text{number of } i\text{'s.}$$

$$\begin{aligned} E(m_i) &= E\left(\sum_{j=1}^m Z_{ij}\right) = \sum_{j=1}^m E(Z_{ij}), \\ &= \sum_{j=1}^m 1 * P(X_j = i) = \sum_{j=1}^m \frac{\pi a_i \theta^i}{f(\theta)}, \\ &= \frac{m \pi a_i \theta^i}{f(\theta)}, \end{aligned}$$

$$\begin{aligned} \text{Now, } E\left(\sum_{i=1}^{\infty} m_i\right) &= \sum_{i=1}^{\infty} E(m_i) = \sum_{i=1}^{\infty} \frac{m \pi a_i \theta^i}{f(\theta)}, \\ &= \frac{m \pi}{f(\theta)} \sum_{i=1}^{\infty} a_i \theta^i = \frac{m \pi}{f(\theta)} (f(\theta) - a_0), \end{aligned}$$

$$\text{Therefore, } E\left(\sum_{i=1}^{\infty} m_i\right) = m \pi \left(1 - \frac{a_0}{f(\theta)}\right).$$

iii. Consider $E\left(\sum_{i=1}^{\infty} i m_i\right) = \sum_{i=1}^{\infty} i E(m_i) = \sum_{i=0}^{\infty} i E(m_i),$

$$\begin{aligned} &= \sum_{i=0}^{\infty} i m P(X = i) = m \pi \sum_{i=0}^{\infty} \frac{i a_i \theta^i}{f(\theta)}, \\ &= \frac{m \pi \theta f'(\theta)}{f(\theta)}. \end{aligned}$$

2. ZERO INFLATED POISSON DISTRIBUTION

A1.2. Let X_1, X_2, \dots, X_m be iid r. v. with ZIPD then,

i. $E(m_0) = m(1 - \pi + \pi e^{-\theta})$

ii $E\left(\sum_{i=1}^{\infty} m_i\right) = m\pi(1 - e^{-\theta})$

iii $E\left(\sum_{i=1}^{\infty} i m_i\right) = m\pi\theta$

Proof.- Define,

$$Y_i = \begin{cases} 1, & \text{if } X_i = 0 \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} E(m_0) &= E\left(\sum_{i=1}^m Y_i\right) \\ &= \sum_{i=1}^m E(Y_i) \\ &= \sum_{i=1}^m 1 P(X_i = 0) \\ &= \sum_{i=1}^m (1 - \pi + \pi e^{-\theta}) \\ &= m(1 - \pi + \pi e^{-\theta}) \end{aligned}$$

ii. To find $E\left(\sum_{i=1}^{\infty} m_i\right)$

Define $Z_{ij} = \begin{cases} 1, & \text{if } X_i = i \\ 0, & \text{otherwise} \end{cases}$

$$\sum_{j=1}^{\infty} Z_{ij} = m_i = \text{number of } i\text{'s.}$$

$$\begin{aligned} E(m_i) &= E\left(\sum_{j=1}^m Z_{ij}\right) = \sum_{j=1}^m E(Z_{ij}), \\ &= \sum_{j=1}^m 1 * P(X_j = i) = \sum_{j=1}^m \frac{\pi e^{-\theta} \theta^i}{i!}, \\ &= \frac{m\pi e^{-\theta} \theta^i}{i!}. \end{aligned}$$

$$\begin{aligned} \text{Now, } E\left(\sum_{i=1}^{\infty} m_i\right) &= \sum_{i=1}^{\infty} E(m_i) \\ &= \sum_{i=1}^{\infty} \frac{m\pi e^{-\theta} \theta^i}{i!}, \\ &= m\pi e^{-\theta} \sum_{i=1}^{\infty} \frac{\theta^i}{i!} \\ &= m\pi(1 - e^{-\theta}). \end{aligned}$$

iii. Consider $E\left(\sum_{i=1}^{\infty} i m_i\right) = \sum_{i=1}^{\infty} i E(m_i)$

$$\begin{aligned} &= \sum_{i=0}^{\infty} i E(m_i) \\ &= \sum_{i=0}^{\infty} i m P(X = i) \\ &= m\pi \sum_{i=0}^{\infty} \frac{i e^{-\theta} \theta^i}{i!} \\ &= m\pi \theta \end{aligned}$$

3. ZERO INFLATED NEGATIVE BINOMIAL DISTRIBUTION

A1.3. Let X_1, X_2, \dots, X_m iid r. v. with ZINBD then,

i. $E(m_0) = m(1 - \pi + \pi Q^{-r}).$

ii $E\left(\sum_{i=1}^{\infty} m_i\right) = m\pi(1 - Q^{-r}).$

iii $E\left(\sum_{i=1}^{\infty} i m_i\right) = m\pi rP.$

Proof.- Define ,

$$Y_i = \begin{cases} 1, & \text{if } X_i = 0 \\ 0, & \text{otherwise} \end{cases}$$

i. $E(m_0) = E\left(\sum_{i=1}^m Y_i\right),$

$$\begin{aligned} &= \sum_{i=1}^m E(Y_i) , \\ &= \sum_{i=1}^m 1 * P(X = 0) , \\ &= \sum_{i=1}^m (1 - \pi + \pi Q^{-r}), \\ &= m(1 - \pi + \pi Q^{-r}). \end{aligned}$$

ii. To find $E\left(\sum_{i=1}^{\infty} m_i\right),$

Define $Z_{ij} = \begin{cases} 1, & \text{if } X_i = i \\ 0, & \text{otherwise} \end{cases}$

$\sum_{j=1}^{\infty} Z_{ij} = m_i = \text{number of } i\text{'s.}$

$$\begin{aligned}
E(m_i) &= E\left(\sum_{j=1}^m Z_{ij}\right) \\
&= \sum_{j=1}^m E(Z_{ij}) \\
&= \sum_{j=1}^m 1 \cdot P(X_j = i) \\
&= \sum_{j=1}^m \pi \binom{r}{i} Q^{-r} \left(\frac{-P}{Q}\right)^i \\
&= m\pi \binom{r}{i} Q^{-r} \left(\frac{-P}{Q}\right)^i
\end{aligned}$$

$$\begin{aligned}
\text{Now, } E\left(\sum_{i=1}^{\infty} m_i\right) &= \sum_{i=1}^{\infty} E(m_i) \\
&= \sum_{i=1}^{\infty} m\pi \binom{r}{i} Q^{-r} \left(\frac{-P}{Q}\right)^i \\
&= m\pi Q^{-r} \sum_{i=1}^{\infty} \binom{r}{i} \left(\frac{-P}{Q}\right)^i \\
&= m\pi (1 - Q^{-r})
\end{aligned}$$

iii. Consider $E\left(\sum_{i=1}^{\infty} i m_i\right) = \sum_{i=1}^{\infty} i E(m_i)$

$$\begin{aligned}
&= \sum_{i=0}^{\infty} i E(m_i) = \sum_{i=0}^{\infty} i m P(X = i) \\
&= m\pi \sum_{i=0}^{\infty} i \binom{r}{i} Q^{-r} \left(\frac{-P}{Q}\right)^i = m\pi r P,
\end{aligned}$$



4. ZERO INFLATED BINOMIAL DISTRIBUTION

A1.4. Let X_1, X_2, \dots, X_m be iid r. v. with ZIBD then,

$$i. \quad E(m_0) = m(1 - \pi + \pi(1 - \theta)^n).$$

$$ii \quad E\left(\sum_{i=1}^n m_i\right) = m\pi(1 - (1 - \theta)^n).$$

$$iii \quad E\left(\sum_{i=1}^n i m_i\right) = mn\pi\theta.$$

Proof.- Define ,

$$Y_k = \begin{cases} 1, & \text{if } X_k = 0, \quad k = 1, 2, \dots, m \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} i. \quad E(m_0) &= E\left(\sum_{k=1}^m Y_k\right) \\ &= \sum_{k=1}^m E(Y_k) = \sum_{k=1}^m 1 * P(X = 0) \\ &= \sum_{k=1}^m (1 - \pi + \pi(1 - \theta)^n) \\ &= m(1 - \pi + \pi(1 - \theta)^n) \end{aligned}$$

$$vii. \quad \text{To find } E\left(\sum_{i=1}^n m_i\right)$$

$$\text{Define ,} \quad Z_{ij} = \begin{cases} 1, & \text{if } X_j = i \quad i = 1, 2, \dots, n, j = 1, 2, \dots, m. \\ 0, & \text{otherwise} \end{cases}$$

$$\sum_{j=1}^m Z_{ij} = m_i = \text{number of } i\text{'s.}$$

$$\begin{aligned}
E(m_i) &= E\left(\sum_{j=1}^m Z_{ij}\right) \\
&= \sum_{j=1}^m E(Z_{ij}) \\
&= \sum_{j=1}^m 1 * P(X_j = i) \\
&= \sum_{j=1}^m \pi \binom{n}{i} \theta^i (1-\theta)^{n-i} \\
&= m \pi \binom{n}{i} \theta^i (1-\theta)^{n-i}
\end{aligned}$$

$$\begin{aligned}
\text{Now, } E\left(\sum_{i=1}^n m_i\right) &= \sum_{i=1}^n E(m_i), \\
&= \sum_{i=1}^n m \pi \binom{n}{i} \theta^i (1-\theta)^{n-i}, \\
&= m \pi (1 - (1-\theta)^n),
\end{aligned}$$

iii. Consider $E\left(\sum_{i=1}^n i m_i\right) = \sum_{i=1}^n i E(m_i)$

$$\begin{aligned}
&= \sum_{i=0}^n i E(m_i) \\
&= \sum_{i=0}^n i m P(X = i) \\
&= m \pi \sum_{i=0}^n i \binom{n}{i} \theta^i (1-\theta)^{n-i} \\
&= m n \pi \theta
\end{aligned}$$